| Question |  | Answer | Marks | Guidance |
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| 1 | (i) | A Normal test is not appropriate since ... ... the sample is small and ... the population variance is not known (it must be estimated from the data). | E1 <br> E1 <br> [2] | Allow use of " $\sigma$ ", otherwise insist on "population". |
| 1 | (ii) | The sample is taken from a Normal population. | B1 [1] |  |
| 1 | (iii) | $\begin{aligned} & \mathrm{H}_{0}: \mu=7.8 \\ & \mathrm{H}_{1}: \mu \neq 7.8 \end{aligned}$ <br> where $\mu$ is the mean water pressure. | B1 | Both hypotheses. Hypotheses in words only must include "population". Do NOT allow " $\bar{X}=\ldots$ " or similar unless $\bar{X}$ is clearly and explicitly stated to be a population mean. <br> For adequate verbal definition. Allow absence of "population" if correct notation $\mu$ is used. |
|  |  | $\bar{x}=7.631 \quad s=0.1547$ <br> Test statistic is $\frac{7.631-7.8}{\frac{0.1547}{\sqrt{9}}}$ | B1 <br> M1 | $s_{\mathrm{n}}=0.1459$ but do $\underline{\text { NOT }}$ allow this here or in construction of test statistic, but ft from there. <br> Allow c's $\bar{X}$ and/or $s_{n-1}$. <br> Allow alternative: $7.8+(c$ 's -2.896$) \times 0.1547 / \sqrt{9}$ (= 7.65...) for subsequent comparison with $\bar{x}$. <br> (Or $\bar{x}-(c$ 's -2.896$) \times 0.1547 / \sqrt{9}(=7.78 \ldots)$ for comparison with 7.8.) |
|  |  | $=-3.27(7)$ | A1 | c.a.o. but ft from here in any case if wrong. Use of $\mu-\bar{x}$ scores M1A0. |
|  |  | Refer to $t_{8}$. <br> Double-tailed 2\% point is $\pm 2.896$. | M1 A1 | No ft from here if wrong. <br> Must compare test statistic with minus 2.896 unless absolute values are being compared. No ft from here if wrong. <br> Allow $\mathrm{P}(t<-3.27(7)$ or $t>3.27(7))=0.0113$ for M1A1. |
|  |  | Significant. <br> Sufficient evidence to suggest that the mean water pressure has changed. | $\begin{aligned} & \text { A1 } \\ & \text { A1 } \end{aligned}$ | ft only c's test statistic if both M's scored. <br> ft only c's test statistic if both M's scored. Conclusion in context to include "average" o.e. |
|  |  |  | [9] |  |


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| 1 | (iv) | In repeated sampling, 95\% of all confidence intervals constructed in this way will contain the true mean. | E1 <br> E1 <br> [2] |  |
| 1 | (v) | CI is given by $7.631 \pm$ $\begin{aligned} & 2 \cdot 306 \\ & \times \frac{0.1547}{\sqrt{9}} \\ &=7.631 \pm 0.118(9)=(7.512,7.750) \end{aligned}$ | M1 <br> B1 <br> M1 <br> A1 <br> [4] | ZERO/4 if not same distribution as test. Same wrong distribution scores maximum M1B0M1A0. Recovery to $t_{8}$ is OK. <br> Allow c's $\bar{x}$. <br> 2.306 seen. <br> Allow c's $s_{n-1}$. <br> c.a.o. Must be expressed as an interval. |
| 2 | (i) |  | G1 <br> G1 <br> G1 <br> [3] | Curve with positive gradient, through the origin and in the first quadrant only. Correct shape for an inverted parabola ending at maximum point. End point (2, 3/4) labelled. |



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| 2 | (iv) | $\begin{aligned} \mathrm{P}(X & <1)=\frac{3}{16} \int_{0}^{1}\left(4 x-x^{2}\right) \mathrm{d} x \\ & =\frac{3}{16}\left[2 x^{2}-\frac{x^{3}}{3}\right]_{0}^{1} \\ & =\frac{3}{16}\left\{\left(2-\frac{1}{3}\right)-0\right\} \\ & =\frac{5}{16} \end{aligned}$ | M1 <br> A1 <br> [2] | Correct integral for $\mathrm{P}(X<1)$ with limits (which may appear later). <br> cao. Condone absence of " -0 " when limits applied. |
| 2 | (v) | Regard the reed beds as clusters. <br> Select a few clusters (maybe only one) at random. <br> Take a (simple random) sample of reeds (or maybe all of them) from the selected cluster(s). | E1 <br> E1 <br> E1 <br> [3] | NB "Clusters of reeds" scores 0 unless clearly and correctly explained. |
| 3 |  | $\begin{aligned} P 1 & \sim \mathrm{~N}\left(2025,44.6^{2}\right) \\ P 2 & \sim \mathrm{~N}\left(1565,21.8^{2}\right) \\ I & \sim \mathrm{~N}\left(1410,33.8^{2}\right) \end{aligned}$ |  | When a candidate's answers suggest that (s)he appears to have neglected to use the difference columns of the Normal distribution tables penalise the first occurrence only. |
| 3 | (i) | $\begin{aligned} & \mathrm{P}(P 1<2100)= \\ & \mathrm{P}\left(Z<\frac{2100-2025}{44.6}\right.=1.681(6)) \\ &=0.9536 / 7 \end{aligned}$ | M1 <br> A1 <br> A1 <br> [3] | For standardising. Award once, here or elsewhere. с.a.o. |


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| 3 | (ii) | $\begin{aligned} & \text { Require } \mathrm{P}(P 1-P 2>400) \\ & P 1-P 2 \sim(2025-1565=460, \\ & \left.44.6^{2}+21.8^{2}=2464.4\right) \end{aligned} \quad \begin{aligned} & \mathrm{P}(\text { this }>400)= \\ & \mathrm{P}\left(Z>\frac{400-460}{\sqrt{2464.4}}=-1.208(6)\right)=0 \cdot 8864 / 5 \end{aligned}$ | M1 <br> B1 <br> B1 <br> A1 <br> [4] | Mean. <br> Variance. Accept sd (= 49.64). <br> cao |
| 3 | (iii) | $\begin{aligned} & T=P 1+P 2+I \sim \mathrm{~N}(5000, \\ & \left.\quad \sigma^{2}=44.6^{2}+21.8^{2}+33.8^{2}=3606.84\right) \\ & \text { Require } b \text { s.t. } \mathrm{P}(T>b)=0.95 \\ & \therefore \frac{b-5000}{\sqrt{3606.84}}=-1.645 \\ & \therefore b=5000-1.645 \times \sqrt{3606.84}=4901.2 . \end{aligned}$ | B1 <br> B1 <br> B1 <br> A1 <br> [4] | Mean. <br> Variance. Accept sd (= 60.056...). <br> -1.645 seen. <br> c.a.o. |
| 3 | (iv) | $\begin{gathered} \text { Mean }=(1.2 \times 2025)+(1.3 \times 1565)+ \\ (0.8 \times 1410)=£ 5592.50 \\ \text { Var }=\left(1.2^{2} \times 44.6^{2}\right)+\left(1.3^{2} \times 21.8^{2}\right)+ \\ \left(0.8^{2} \times 33.8^{2}\right)=4398.7076 \approx £^{2} 4399 \end{gathered}$ | $\begin{aligned} & \text { B1 } \\ & \text { M1 } \\ & \text { A1 } \\ & \text { [3] } \end{aligned}$ | Condone absence of $£$. <br> Use of at least one of $\left(1.2^{2} \times 44.6^{2}\right)$ etc... Condone absence of $£^{2}$. |
| 3 | (v) | $\begin{aligned} & \text { Mean }=(123.72+127.38) / 2=125.55 \\ & s=\frac{127.38-125.55}{2.576 / \sqrt{50}}=5.02(3) \end{aligned}$ | $\begin{gathered} \text { B1 } \\ \text { B1 } \\ \text { M1 } \\ \text { A1 } \\ {[4]} \\ \hline \end{gathered}$ | Cao <br> Sight of 2.576 . <br> Or equivalent. cao |




